# Verifying Differentially Private Bayesian Inference



Joint work with G. Barthe, G.P. Farina, E.J. Gallego Arias, A.Gordon,...

#### Differentially Private vs Probabilistic Inference

Differential Privacy Probabilistic Inference

#### Differentially Private vs Probabilistic Inference

Differential Privacy Probabilistic Inference

The goal in machine learning is very often similar to the goal in private data analysis. The learner typically wishes to learn some simple rule that explains a data set. However, she wishes this rule to generalize [...] Generally, this means that she wants to learn a rule that captures distributional information about the data set on hand, in a way that does not depend too specifically on any single data point.

C. Dwork & A. Roth - The algorithmic Foundations of Differential privacy

# Outline

- Bayesian learning
- A language for bayesian learning
- Differential Privacy
- Combining differential privacy and bayesian learning

# Bayesian Inference

# Probabilistic Inference



# Bayes theorem





# A simple example: bias of a coin





# Model/Likelihood: Bernoulli

**Observed values** 

 $\mathbf{y} = (0, 1, 0, 1, 1, 1, ..., 0)$ 

# Model/Likelihood: Bernoulli

**Observed values** 

 $\mathbf{y} = (0, 1, 0, 1, 1, 1, ..., 0)$ 

The probability mass function  $f(s, \Theta)$  is

$$f(s,\Theta) = \begin{cases} \Theta & \text{if } s = 1 \\ 1 - \Theta & \text{if } s = 0 \end{cases}$$

where  $\Theta$  is the "bias of the coin". This can also be expressed as:

$$f(s,\Theta) = \Theta^{s}(1-\Theta)^{1-s}$$

# Prior: Beta distribution

**Probability Distribution Function** 



Conjugate to the Bernoulli distribution

## Prior: Beta distribution



# Bias of a Coin

Beta(1500,1500)







# Another example



Known variance

# Another example



# Graphical Models

	I	disease
	1	0
1	2	1
4	3	0
4	4	1
Į	5	1
(	6	1
	7	0
ł	8	1
(	9	





# Probabilistic PCF

$$\begin{array}{lll} e \in \mathbf{PCF}_p(\mathcal{X}) & ::= & x \mid c \mid e \mid e \mid \lambda x. e \mid \mathsf{let} \ x = e \ \mathsf{in} \ e \\ & \mid & \mathsf{letrec} \ f \ x = e \mid \mathsf{case} \ e \ \mathsf{with} \ [d_i \ \overline{x_i} \Rightarrow e_i]_i \\ & \mid & \mathsf{unit}_{\mathsf{M}} \ e \mid \mathsf{bind}_{\mathsf{M}} \ x = e \ \mathsf{in} \ e \\ & \mid & \mathsf{observe} \ x \Rightarrow e \ \mathsf{in} \ e \mid \mathsf{infer}[e] \mid \mathsf{ran}(e) \\ & \mid & \mathsf{bernoulli}(e) \mid \mathsf{gauss}(e, e) \mid \mathsf{beta}(e, e) \mid \mathsf{uniform}() \\ & \mid & \mathsf{laplace}(e, e) \mid \mathsf{exponential}(e, e, e) \end{array}$$

# Probabilistic Semantics

$$\begin{split} \llbracket \Gamma \vdash \texttt{unit}_{M} e : \mathfrak{M}[\tau] \rrbracket_{\theta} = x \mapsto \begin{cases} 1 \text{ if } x = \llbracket e \rrbracket_{\theta} \\ 0 \text{ otherwise} \end{cases} \\ \llbracket \Gamma \vdash \texttt{bind}_{M} x = e_{1} \text{ in } e_{2} : \mathfrak{M}[\sigma] \rrbracket_{\theta} = \\ d \mapsto \sum_{g \in \llbracket \tau \rrbracket} (\llbracket e_{1} \rrbracket_{\theta}(g) \times \llbracket e_{2} \rrbracket_{\theta}^{g}(d)) \\ \llbracket \Gamma \vdash \texttt{letrec}^{\mathfrak{m}} f x = e : \tau \to \sigma \rrbracket_{\theta} = \bigcup_{n \in \mathbb{N}} F^{n}(\bot_{\llbracket \tau \to \sigma} \rrbracket) \\ \texttt{where } F(d) = \llbracket \Gamma, f : \tau \to \sigma \vdash \lambda x. e : \tau \to \sigma \rrbracket_{\theta}^{d}_{f} \end{split}$$

# Observe for Conditional Distributions



# Observe for Conditional Distributions



$$P(y|x) = \frac{P(x|y)P(y)}{P(x)}$$

# Semantics of Observe

 $\llbracket \Gamma \vdash \mathsf{observe} \, x \Rightarrow t \, \mathsf{in} \, u : \mathfrak{M}(A) \rrbracket_{\rho} =$ 

$$d \mapsto \frac{\llbracket \Gamma \vdash u : \mathfrak{M}[A] \rrbracket_{\rho} d \times (\llbracket \Gamma, x : A \vdash t : \mathsf{bool} \rrbracket_{\rho_x^d} \mathsf{true})}{\sum_{a \in \llbracket A \rrbracket} \llbracket \Gamma \vdash u : \mathfrak{M}[A] \rrbracket_{\rho} a \times (\llbracket \Gamma, x : A \vdash t : \mathsf{bool} \rrbracket_{\rho_x^a} \mathsf{true})}$$

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## Filter and rescaling

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## Filter and rescaling







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# $(\epsilon, \delta)$ -Differential Privacy

# $\begin{array}{l} \textbf{Definition} \\ \textbf{Given } \boldsymbol{\epsilon}, \boldsymbol{\delta} \geq 0, \ a \ probabilistic \ query \ Q: db \rightarrow R \ is \\ (\boldsymbol{\epsilon}, \boldsymbol{\delta}) \text{-differentially private iff} \\ \forall b_1, b_2: db \ differing \ in \ one \ row \ and \ for \ every \ S \subseteq R: \\ Pr[Q(b_1) \in S] \leq exp(\boldsymbol{\epsilon}) \cdot Pr[Q(b_2) \in S] + \boldsymbol{\delta} \end{array}$
# $(\varepsilon, \delta)$ -Differential Privacy

A query returning a probability distribution

 $\begin{array}{l} \text{Definition} \\ \text{Given } \epsilon, \delta \geq 0, \text{ a probabilistic query } Q: db \rightarrow R \text{ is} \\ (\epsilon, \delta) \text{-differentially private iff} \\ \forall b_1, b_2: db \text{ differing in one row and for every } S \subseteq R: \\ Pr[Q(b_1) \in S] \leq \exp(\epsilon) \cdot Pr[Q(b_2) \in S] + \delta \end{array}$ 



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a quantification over all the databases

# $(\epsilon, \delta)$ -Differential Privacy



# $(\epsilon, \delta)$ -Differential Privacy

#### **Definition** Given $\varepsilon, \delta \ge 0$ , a probabilistic query Q: db $\rightarrow R$ is $(\varepsilon, \delta)$ -differentially private iff $\forall b_1, b_2$ :db differing in one row and for every $S \subseteq R$ : $Pr[Q(b_1) \in S] \le exp(\varepsilon) \cdot Pr[Q(b_2) \in S] + \delta^{\uparrow}$

and over all the possible outcomes

# ε-Differential Privacy

#### Definition

Given  $\varepsilon \ge 0$ , a probabilistic query Q: db  $\rightarrow$  R is  $\varepsilon$ -differentially private iff  $\forall b_1$ ,  $b_2$ : db differing in one row and for every S  $\subseteq$  R: Pr[Q(b\_1)  $\in$  S]  $\le \exp(\varepsilon) \cdot \Pr[Q(b_2) \in S]$ 

# ε-Differential Privacy

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Given  $\varepsilon \ge 0$ , a probabilistic query Q: db  $\rightarrow$  R is  $\varepsilon$ -differentially private iff  $\forall b_1$ ,  $b_2$ : db differing in one row and for every S  $\subseteq$  R: Pr[Q(b\_1)  $\in$  S]  $\le \exp(\varepsilon) \cdot \Pr[Q(b_2) \in S]$ 

Let's substitute a concrete instance:  $Pr[Q(b \cup \{x\}) \in S] \leq exp(\epsilon) \cdot Pr[Q(b) \in S]$ Let's use the two quantifiers:  $exp(-\epsilon) \cdot Pr[Q(b) \in S] \leq Pr[Q(b \cup \{x\}) \in S] \leq exp(\epsilon) \cdot Pr[Q(b) \in S]$ and so:

$$\log \left| \frac{\Pr[Q(b \cup \{x\}) \in S]}{\Pr[Q(b) \in S]} \right| \leq \epsilon$$









# Probability of a bad event



# Probability of a bad event



# $(\varepsilon, \delta)$ -Differential privacy

#### for $b \sim_1 b'$ $\Pr[Q(b') \in S] \leq e^{\epsilon} \Pr[Q(b) \in S] + \delta$



							$\bigwedge$	Noise
			19144	02-15-1964	flue	Į		<u> </u>
			19146	05-22-1955	brain tumor	3	sorry	
144	02-	15-1964	34505	5 11-01-1988	depression	ΙŽ	()	
<u> </u>	02	13-1304	25012	2 03-12-1972	diabets	ξ.	5	2
146	05-	22-1955	16544	06-14-1956	anemia	<u>}</u>	A TRO	2
505	11-	01-1988	i I C	pression	17		5 1	
010					$\rightarrow$		51/1	
012	-60	12-19/2		liabets			5	
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			۲			all	i i Ogi ai	<u> </u>

# Sensitivity





# Sensitivity



# Calibrating the Noise on the Sensitivity



#### Components

- Relational Refinement Types
- Higher Order Refinements
- Partiality Monad
- Semantic Subtyping
- Approximate Equivalence for Distributions
  Barthe et al, POPL'15

$$\begin{array}{rll} e \in \mathbf{PCF}_p(\mathcal{X}) & ::= & x \mid c \mid e \; e \mid \lambda x. \, e \mid \texttt{let} \; x = e \; \texttt{in} \; e \\ & \mid & \texttt{letrec} \; f \; x = e \mid \texttt{case} \; e \; \texttt{with} \; [d_i \; \overline{x_i} \Rightarrow e_i]_i \\ & \mid & \texttt{unit}_{\mathrm{M}} \; e \mid \texttt{bind}_{\mathrm{M}} \; x = e \; \texttt{in} \; e \end{array}$$

#### Components

Relational Refinement Types

- Higher Order Refinements
- Partiality Monad
- Semantic Subtyping
- Approximate Equivalence for Distributions Barthe et al, POPI'15

reasoning about two runs of a program

#### Program Logic



#### Program Logic



#### Program Logic



#### Refinement Type



#### Refinement Type



#### Refinement Type



#### Example exp: $\{x \mid x \ge 0\} \rightarrow \{y \mid y \ge 1\}$

## Differential privacy as a Relational Property



## Differential privacy as a Relational Property



# HOARe<sup>2</sup>: Relational Refinement Types

 $P: \{x \mid Pre(x_1, x_2)\} \rightarrow \{y \mid Post(y_1, y_2)\}$ 

Program P Precondition - Postcondition

# HOARe<sup>2</sup>: Relational Refinement Types



# HOARe<sup>2</sup>: Relational Refinement Types



#### Example: Monotonicity of exponential $exp: \{x \mid x_1 \le x_2\} \rightarrow \{y \mid y_1 \le y_2\}$

#### Components

- Relational Refinement Types
- Higher Order Refinements
- Partiality Monad
- Semantic Subtyping

reasoning about DP

Approximate Equivalence for Distributions

Barthe et al, POPL'15



# Lifting of P

• Given dist.  $\mu_1$  over A and  $\mu_2$  over B:

#### $\mu_1 P^* \mu_2$

iff it exists a dist.  $\mu$  over AxB s.t.

- $\mu(x) > 0$  implies  $x \in R$
- $\pi_{1}\mu \leq \mu_{1}$  and  $\pi_{2}\mu \leq \mu_{2}$

# $(\epsilon, \delta)$ -Lifting of P

• Given dist.  $\mu_1$  over A and  $\mu_2$  over B:

 $\mu_1 \ P_{\epsilon\delta}^* \ \mu_2$ 

- iff it exists a dist.  $\mu$  over AxB s.t.
- $\mu(x) > 0$  implies  $x \in R$
- $\pi_{1}\mu \leq \mu_{1}$  and  $\pi_{2}\mu \leq \mu_{2}$
- $\max_{A}(\pi_{i}\mu(A) e^{\varepsilon}\mu_{i}(A), \mu_{i}(A) e^{\varepsilon}\pi_{i}\mu(A)) \leq \delta$

# Verifying Differential Privacy

If we can conclude

$$C: \{x:db \mid d(y_1, y_2) \leq 1\} \rightarrow \{y:O \mid y_1 = y_2^*, y_2\}$$

then C is  $(\varepsilon, \delta)$ -differentially private.

#### Programming Languages for Differentially Private Probabilistic Inference

Differential Privacy Probabilistic Inference
#### Programming Languages for Differentially Private Probabilistic Inference



























## Adding noise on the data



```
function privBerInput (l: B list) (p1: R) (p2: R): M[(0,1)]{
  let function vExp (l: B list) : M[B list]{
    match 1 with
       |nil -> mreturn nil
       |x::xs -> coercion
                   (\exp eps((0,0) \rightarrow 1, (0,1) \rightarrow 0, (1,1) \rightarrow 1,
                             (1, 0) \rightarrow 0 x) :: (vExp 1)
  } in
  mlet nl = (vExp l) in
  let prior = mreturn(beta(p1, p2)) in
  let function Ber (l: B list) (p:M[(0,1)]): M[(0,1)]{
    match 1 with
       |nil -> ran(p)
       |x::xs \rightarrow observe y => y = x in (Ber xs p)
  }
  in
  mreturn(infer (Ber nl prior))
}
```

```
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### Adding noise on the output







Releasing the Parameters



Releasing the Parameters

Sampling from the Distribution



#### Releasing the Parameters

Sampling from the Distribution

```
function privBerInput (l: B list) (p1: R) (p2: R): M[(0,1)]
  let function hellingerDistance (a0:R) (b0:R) (a1:R) (b1:R) : R {
        let gamma (r:R) = (r-1)! in
        let betaf (a:R) (b:R) = gamma(a) * gamma(b) ) / gamma(a+b) in
        let num=betaf ((a0+a1)/2.0) ((b0+b1)/2.0) in
        let denum=Math.Sqrt((betaf a0 b0)*(betaf a1 b1)) in
        Math.Sqrt(1.0-(num/denum))
  }
  in
  let function score (input:M[(0,1)]) (output:M[(0,1)]) : R {
    let beta(a0, b0) = input in
    let beta(a1, b1) = output in
    (-1.0) * (hellingerDistance a0 b0 a1 b1)
  } in
  let prior = mreturn(beta(p1, p2)) in
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```

A distance over distributions

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  }
```

exp eps score (infer (Ber l prior))

in



A distance over distributions

# More general Lifting of P



## Accuracy

Different ways of adding noise can have different accuracy.

- we have theoretical accuracy (how to integrate it in a framework for reasoning about DP?)
- we have experimental accuracy (we need a framework for for test our programs)



#### Tänan teid väga!